

Competitive state of movements during planning predicts sequence performance

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Journal of Neurophysiology

DOI: https://doi.org/10.1152/jn.00645.2020

Published: 01/04/2021

Peer reviewed version

Cyswllt i'r cyhoeddiad / Link to publication

Dyfyniad o'r fersiwn a gyhoeddwyd / Citation for published version (APA): Mantziara, M., Ivanov, T., Houghton, G., & Kornysheva, K. (2021). Competitive state of movements during planning predicts sequence performance. *Journal of Neurophysiology*, *125*(4), 1251-1268. https://doi.org/10.1152/jn.00645.2020

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1	Competitive state of movements during planning
2	predicts sequence performance
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12	Word count: New & Noteworthy: 71; Abstract: 207; Introduction: 1312; Discussion:
13	2243; Conclusions: 179; Methods: 3555; Results: 5002; 49 pages, 5 main figures; 6
14	supplemental figures; 4 supplemental tables.
15	
16	Author contributions: M.M. and K.K. conceived the experiments; M.M., G.H., and
17	K.K. formulated the hypothesis; M.M. and T.I. collected the data; M.M., G.H., and K.K.
18	designed the analysis; M.M., T.I. and K.K. performed the analysis; M.M., T.I., G.H. and
19	K.K. wrote the paper. All authors contributed to editing of the manuscript.
20	
21	Acknowledgements: The authors wish to thank Tom Hartley, Ken Valyear and Simon
22	Watt for helpful comments on the study, and Willem Verwey for useful feedback on an
23	earlier version of the manuscript.
24	
25	Disclosures: The authors declare no conflicts of interest.

26 New & Noteworthy

Sequence planning is an integral part of motor sequence control. Here, we demonstrate that the competitive state of sequential movements during sequence planning can be read out behaviourally through movement probes. We show that position-dependent differences in movement availability during planning reflect sequence preparedness and skill, but not the timing of the planned sequence. Behavioural access to the preparatory state of movements may serve as a marker of sequence planning capacity.

34 Abstract

35 Humans can learn and produce skilled movement sequences from memory, yet 36 the nature of sequence planning is not well understood. Previous computational and 37 neurophysiological work suggests that movements in a sequence are planned as parallel graded activations and selected for output through competition. However, the 38 39 relevance of this planning pattern to sequence production fluency and accuracy, as opposed to the temporal structure of sequences, is unclear. To resolve this question, 40 41 we assessed the relative availability of constituent movements behaviourally during 42 the preparation of motor sequences from memory. In three separate multi-session 43 experiments, healthy participants were trained to retrieve and produce 4-element 44 finger press sequences with particular timing according to an abstract sequence cue. 45 We evaluated reaction time (RT) and error rate as markers of movement availability to 46 constituent movement probes. Our results demonstrate that longer preparation time 47 produces more pronounced differences in availability between adjacent sequence 48 elements, whilst no effect was found for sequence speed or temporal grouping. 49 Further, participants with larger position-dependent differences in movement 50 availability tended to initiate correct sequences faster and with a higher temporal 51 accuracy. Our results suggest that competitive pre-activation during sequence 52 planning is established gradually during sequence planning and predicts sequence skill, rather than the temporal structure of the motor sequence. 53

54

55 **Keywords:**

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motor planning; sequence control; competitive queuing; reaction time; error rate

57 Introduction

58 Producing movement sequences from memory fluently is an essential capacity 59 of primates, in particular humans. It enables a skilled and flexible interaction with the world for a range of everyday activities - from tool-use, speech, and gestural 60 61 communication, to sports and music. Key to fluent sequence production is sequence 62 planning before the initiation of the first movement (Lashley 1951; Rosenbaum 1985), with longer preparation time benefitting sequence execution, i.e., reducing initiation 63 time after a Go cue and improving accuracy (Ariani and Diedrichsen 2019). However, 64 the underlying nature and content of sequence planning is still debated (Remington et 65 66 al. 2018).

67 Different computational accounts of sequence control make contrasting 68 predictions with regard to the content of sequence planning. Models postulating a 69 purely serial control of motor sequences suggest that a well-learnt sequence is a 70 cohesive entity, rather than a series of individual movements, e.g. individual strokes 71 when drawing a geometrical figure or finger presses playing the piano (Goudar and 72 Buonomano 2018; Laje and Buonomano 2013). They predict that sequence planning 73 activity reflects bringing the neural trajectory towards the correct neural state of 74 sequence initiation from which it cascades serially through a learnt trajectory. 75 Sequence planning would therefore entail the preparation of the state occupied by the 76 first movement, e.g. using a null-state to allow preparation without premature initiation, 77 as shown empirically for reaching movements (Kaufman et al. 2014; O'Shea and 78 Shenoy 2016).

79 In contrast, models postulating parallel sequence control, such as competitive 80 queuing models (Houghton 1990), propose simultaneous control of the items, here 81 movements, in a sequence. They predict that preparatory neural activity pre-activates 82 sequence movements *concurrently*. Specifically, the neural activation pattern for each movement is weighted according to its temporal position in the respective sequence 83 84 (Burgess and Hitch 1999; Hartley and Houghton 1996), resulting in a positiondependent pre-activation gradient for each upcoming movement in the sequence. 85 Indirect support for parallel and independent neural control of sequential movements 86 87 stems from observations of serial recall including transposition of neighbouring 88 sequence items and items occupying the same position in different chunks (Glasspool

91

and Houghton 2005; Hartley and Houghton 1996; Henson 1998), and excitability of
forthcoming movements during sequence production (Behmer et al. 2018).

Direct neurophysiological support for the parallel control of sequence

- 92 movements has been provided in the context of well-trained finger sequences (Kornysheva et al. 2019; Pinet et al. 2019), saccades (Basu and Murthy 2020), 93 94 drawing of geometrical shapes (Averbeck et al. 2002). Specifically, during planning, 95 the probability of neural patterns associated with each movement in the sequence was 96 highest for the first, and lowest for the fourth and fifth movements of the planned 97 sequence. This effect could not be explained by a graded pre-pressing of the 98 corresponding fingers according to their order and was observed at the trial-by-trial 99 level, suggesting that this competitive pre-activation is not an artefact of trial averaging 100 (Kornysheva et al. 2019). Importantly, the ordered pre-activation gradient of sequence 101 movements during planning was relevant to subsequent execution. In particular, the 102 quality and strength of this gradient was predictive of sequence production accuracy 103 such that participants with stronger pre-activation differences between the sequence 104 items during planning were more accurate during sequence production. Together, 105 these data suggest that skilled sequence production involves an orderly parallel 106 planning of several movements in advance before sequence initiation and predicts
- 107 better sequence performance.
- 108 While the pre-activation gradient during planning has been shown to predict 109 subsequent execution, it remains unclear what this preparatory pattern reflects – the 110 skill of sequence production (fluency of initiation and accuracy of the sequence 111 execution), or the temporal structure of the sequence (speed and temporal grouping). 112 Most competitive queuing models assume the presence of a temporal or positional 113 context layer and that the activity gradients are learned by associations of the latter to 114 each sequence item in the parallel planning layer, e.g. through Hebbian learning (Burgess and Hitch 1999). The form of activity in the context layer can be as simple 115 116 as a decaying start signal (Page and Norris 1998), a combination of start and end 117 signals (Houghton 1990, 2018) or a sequence of overlapping states (Burgess and 118 Hitch 1999, 2006). Although primarily encoding serial order of sequence items, models 119 utilizing overlapping states can implement effects of temporal grouping or sequence 120 rhythm (Burgess and Hitch 1999; Hartley et al. 2016) making timing an intrinsic

121 property of the competitive queuing of sequential movements. Likewise, a separate 122 timing process (Kornysheva et al. 2013; Kornysheva and Diedrichsen 2014; Medina 123 et al. 2005; Spencer et al. 2009; Ullén and Bengtsson 2003; Zeid and Bullock 2019) 124 may modulate the parallel planning of the serial order of items, e.g. in the parallel 125 planning layer. In both cases, the competitive pre-activation gradient of movements 126 during planning would reflect the temporal grouping or temporal proximity of 127 movements in the upcoming sequence, with movements closer together in time having 128 more similar levels of pre-activation than those that are further apart (Burgess and 129 Hitch 1999). In contrast, sequence timing may not impact the competitive pre-130 activation of sequential movements during planning and interact with the latter during 131 execution only.

132 In order to investigate the nature of sequence planning and its relation to 133 subsequent execution, we developed a behavioural paradigm to capture the 134 preparatory state of each constituent movement of a well-learned sequence during 135 planning. Following training, participants prepared a motor sequence from memory 136 following an abstract visual stimulus associated with a particular sequence of finger 137 presses performed with a particular speed or temporal grouping. In half of the trials 138 during the test phase, the Go cue was replaced by a finger press cue prompting the 139 production of movements associated with different positions in the sequence. We used 140 behavioural availability for fast and correct execution of the presses in these Probe 141 trials (RT and error rate) as behavioural markers of the relative pre-activation of 142 upcoming movements during sequence planning.

143 We hypothesized that if competitive queuing during planning primarily reflected 144 the accuracy of the sequence plan (Averbeck et al. 2002; Kornysheva et al. 2019), but 145 not its timing, we would predict a gradual differentiation of the position-dependent pre-146 activation gradient with longer sequence preparation time. Accordingly, we would 147 observe an increase of position-dependent differences in press availability across 148 preparation durations of 500, 1000 and 1500 ms, despite matched speed and temporal 149 grouping of sequence production. Further, participants with a more pronounced 150 gradient would be more fluent and accurate, specifically show more rapid sequence 151 initiation of correct sequences after the Go cue, more accurate timing and fewer finger 152 press errors.

Alternatively, if the gradient reflected the timing of the sequence during 153 154 planning, movements planned to be executed closer in time would show smaller 155 position-dependent differences relative to movements further apart. Accordingly, 156 sequences twice as fast (speed manipulation) would result in more similar levels of 157 availability of movements in neighbouring sequence positions. Further, the latter would 158 be modulated by irregular inter-press-intervals (IPI) with shorter versus longer IPIs 159 being accompanied by smaller versus larger differences in position-dependent 160 availability during planning, respectively (temporal grouping manipulation).

161 We report that during the 1.5 seconds of sequence retrieval and preparation 162 from memory the behavioural availability of sequential movements decreases on 163 average with their planned serial position, up to the last but one. Specifically, 164 movement probes associated with later sequence positions were progressively more 165 likely to lead to erroneous presses during planning, and correct presses were executed 166 more slowly. This characteristic preparatory gradient of movement availability 167 increased with preparation duration rendering movements pre-planned to occur in later 168 compared to earlier sequence positions progressively less available. Across 169 participants, the size of this gradient during preparation correlated with more fluent 170 initiation and temporally accurate sequence production. Contrary to the timing 171 hypothesis, we found no reliable effect of sequence speed or temporal grouping on 172 movement availability during planning. Based on this data, we propose that sequence 173 planning involves a competitive pre-activation gradient of sequential movements 174 during sequence planning which operates independently of sequence timing and 175 facilitates skilled sequence performance.

176

177 Materials and Methods

178 *Participants*

Data were collected from a total of 55 right-handed University students (Experiment 1: N = 19, 11 females; M = 24.2 years, SD = 4.1; Experiment 2: N=18, 11 females; M = 24.2 years, SD = 4.5; Experiment 3: N = 18, 9 females; M = 20.8 years, SD = 2.4). Four additional participants were tested but excluded from analysis based on their sequence production finger error rate (cf. Participant exclusion criteria). They were hypothesis-naive and had no previous exposure in performing a similar

experimental task. All participants had normal or corrected-to-normal vision and 185 186 reported no history of neurological or psychiatric disorders or hearing problems. 187 Handedness was evaluated through the online Handedness Questionnaire 188 (http://www.brainmapping.org/shared/Edinburgh.php) adapted from the Edinburgh Handedness Inventory (Oldfield 1971) (Experiment 1, *M* = 88.4, *SD* = 9.4; Experiment 189 2, M = 90.6, SD = 9.7; Experiment 3, M = 90, SD = 11.8). All participants provided 190 191 written informed consent before participation and were debriefed after completing the 192 study. They were compensated either monetarily or with course credits at the end of 193 the experiment. All procedures were approved by the Bangor University School of 194 Psychology Research Ethics Committee (Ethics Review Board Approval Code 2017-195 16100-A14320).

196 Apparatus

197 For all three experiments participants were seated in a guiet room in front of a 198 19-inch LCD monitor (LG Flatron L1953HR, 1280 x 1024 pixels), wearing headphones 199 for noise isolation. All instructions, visual stimuli and feedback were precisely timed by 200 the monitor's refresh rate (60Hz) and controlled by Cogent 2000 (v1.29) 201 (http://www.vislab.ucl.ac.uk/cogent.php) through a custom-written MATLAB program 202 (v9.2 R2017a, The MathWorks, Inc., Natick, Massachusetts, United States). In 203 Experiments 1 and 2, a Pyka 5-button fiber optic device (Current Designs) was used 204 to record the responses. A customized foam channel stabilized the cable and a thin anti-slip mat, placed underneath the response device, prevented from sliding during 205 206 the task. The response device was positioned horizontally and adjusted for each 207 participant to ensure good control over the target buttons as well as arm and wrist 208 comfort. Participants were instructed to place the right index, middle, ring and little 209 fingers on the respective target buttons of the device. Experiment 3 used an identical 210 experimental set-up with the exception that responses were recorded using a 211 computer keyboard. Here, participants were instructed to place their right thumb in 212 addition to the rest of the right-hand fingers on the designated keyboard keys. For 213 hand stabilization and comfort their wrist was positioned on a rest cushion.

214

Experimental design

All three experiments employed a visually cued motor learning task adapted from Kornysheva et al. (2019). Experiments 1 and 2 involved the recording of

sequential and single button presses produced with the four fingers (index, middle, 217 218 ring and little) of the right hand. Experiment 3 additionally required single presses with 219 the thumb. In all experiments, participants were trained to associate a visual cue (an 220 abstract fractal shape, henceforth Sequence cue) with a four-element finger sequence 221 produced with a specific timing. The paradigm employed two main trial types: 222 Sequence and Probe (single press) trials. Sequence trials were further divided into 223 visually instructed and memory-guided trials. Instructed Sequence trials involved the 224 presentation of four visual digit cues (index, middle, ring and little) at specified intervals 225 comprising a unique target sequence. These were only used during training in the first 226 two days, and during two refresher blocks on the third day (Figure 1a). The test phase 227 on the third day involved sequence production without visual guidance (Figure 1b). 228 *Probe* trials involved the production of only one visual digit cue (*Probe* cue) 229 corresponding to one of the serial positions in the target sequence (Figure 1c).

Experiment 1 – Preparation duration. All participants were trained to produce two different finger sequences comprising four presses with target IPIs of 800 ms (slow timing). Two additional sequences served as practice sequences to impose familiarization with the task. All sequences were randomly generated offline for each participant through a custom-written MATLAB code for each participant. The sequence generation process excluded sequences with ascending and descending digit triplets and identical finger positions.

All trial types started with a *Sequence* cue. The *Sequence* cue had a fixed duration of 400 ms followed by a fixation cross, the latency of which varied depending on the delay period between the *Sequence* cue and *Go* cue onsets. The resultant short (500 ms), intermediate (1000 ms), and long (1500 ms) delay periods comprised the three preparation duration conditions employed in the task. After the delay period, a black right-hand stimulus appeared as the *Go* cue.

In an instructed *Sequence* trial, the *Go* cue was presented on a grey background for 2400 ms. A white circle appeared on top of the corresponding finger digits of the hand stimulus sequentially to guide the participants throughout the execution of the sequence. The time intervals between the digit cues formed the target timing of the sequence and defined its duration of 2400 ms. To achieve finger and temporal accuracy during training, participants were asked to "synchronise" the correct

finger presses with the digit cues until the completion of the sequence. As the first digit 249 250 cue of a sequence appeared at the same time as the Go cue, immediate initiation of 251 the sequence was emphasized in the instructions. In a memory-guided Sequence trial, 252 the Go cue was presented on a green background, remaining on the screen for 2400 253 ms. Memory-guided Sequence trials were devoid of finger digit cues, requiring 254 participants to produce the upcoming target sequence from memory. Participants were 255 instructed to initiate the sequence as quickly as possible and produce the sequence 256 according to its target finger order and timing. In a *Probe* trial, after the delay period, 257 the Go cue was replaced with a Probe cue, namely a single digit cue, displayed for 258 1000 ms. The *Probe* cue prompted a single press with a corresponding finger as fast 259 and accurately as possible. Participants were encouraged to avoid premature 260 responses (before the Go cue) in all trial types. Following the Go cue in any trial type, 261 a fixation cross (1000 ms) and, subsequently, feedback (1000 ms) were presented on 262 the screen. The duration of a Sequence trial was 5.4 s, while a Probe trial had a 263 duration of 4 s, including feedback. The inter-trial-interval (ITI) was fixed at 800 ms.

264 The experiment consisted of two 90 min long training sessions (Days 1 and 2) 265 and a test session (Day 3) which took place over three consecutive days. Day 1 266 commenced with a practice block which involved two instructed and two memory-267 guided Sequence trials for each of the target sequences, and two randomly selected 268 *Probe* trials, with randomly chosen preparation durations. Over the three days, 269 participants serially underwent a pre-training (2 blocks), a training (36 blocks), a post-270 training (2 blocks) and a test phase (2 refresher training blocks + 16 test blocks), 271 completing a total of 58 blocks. To assess sequence planning and execution from 272 memory only data from the test phase is presented here.

273 Participants were naïve as to the structure of the transition from the training 274 through to the test phase and which block type they were administered (Figure 1a). 275 The training phase was organized in three stages: 12 blocks of 288 instructed 276 Sequence and 72 Probe trials (stage A, 80% instructed Sequence and 20% Probe 277 trials in each block), 12 blocks of 144 instructed, 144 memory-guided Sequence and 278 72 Probe trials (stage B, 40% for each Sequence type and 20% Probe trials in each 279 block), and 12 blocks of 288 memory-guided Sequence and 72 Probe trials (stage C, 280 80% memory-guided Sequence and 20% Probe trials in each block). Each training

block (3 min long) consisted of 30 trials. On each training block there was a 20% 281 282 occurrence of *Probe* trials (6 in each block) comprising a total of 216 throughout the 283 training blocks. All Probe trial conditions (24; 2 sequences × 3 preparation durations × 284 4 digits) were counterbalanced across the training blocks. The test phase (Day 3) started with two refresher training blocks (stage B, 40% for each Sequence type and 285 286 20% *Probe* trials in each block) and immediately progressed to 16 blocks of 48 trials 287 each, in which 24 memory-quided Sequence and 24 Probe trials were randomly 288 presented (test, 50% memory-guided Sequence and 50% Probe trials). Duration of 289 each test block was 4.4 min. The preparation duration conditions were 290 counterbalanced across the two target sequences in memory-guided Sequence and 291 *Probe* trials in each block. This resulted in a total of 128 memory-guided Sequence 292 trials per preparation duration condition, across blocks. In Probe trials, each Probe cue 293 was combined with the three preparation duration conditions resulting in 32 trials per 294 digit cue per preparation duration condition. The test phase had a total of 768 trials 295 (384 memory-guided Sequence and 384 Probe trials). Overall, the participants 296 underwent 2004 trials excluding the practice trials.

297 Preparation duration (foreperiod) effects on RT have been associated with 298 carry-over effects from preceding to current trials, and may bias our RT findings if trial 299 history is unbalanced (Langner et al. 2018; Steinborn and Langner 2012). Post-hoc, 300 we examined the preparation duration conditions in both Probe trials and memory-301 guided Sequence trials (cf. Supplemental Figure S1a, b). The mean preparation 302 duration of preceding trials (previous, *n*-1, or two trials previously, *n*-2) did not vary 303 depending on the serial position associated with the target sequence in any of the 304 preparation durations of a current trial (n) (4 x 3 repeated-measures ANOVAs: Position 305 x Preparation duration n-1, F(6, 108) = .88, p = .511, $\eta p^2 = .05$; Position x Preparation duration *n*-2, *F* (6, 108) = 1.14, p = .344, $np^2 = .06$). Equally, analysis of the sequence 306 307 production trials revealed that preparation duration of a current trial did not vary with 308 the mean preparation duration of preceding trials (one-way repeated-measures 309 ANOVAs: Preparation duration *n*-1, *F* (2, 36) = 2.53, *p* = .093, ηp^2 = .12; Preparation duration n-2, F(2, 36) = .36, p = .701, $\eta p^2 = .02$). This demonstrates a balanced design 310 311 in which the foreperiod length history up to two previous trials was unlikely to bias RT 312 or error rates on the current trial.

Experiment 2 – Sequence timing. Procedures for Experiment 2 were identical 313 314 to Experiment 1 except that the delay period was fixed at 1500 ms and participants 315 were trained in associating three target sequences. Each featured a unique Sequence 316 cue associated with one finger order instructed to be performed at three target IPIs: 317 slow (800-800-800 ms), fast (400-400-400 ms) and irregular (400-1600-400 ms), 318 comprising the three timing conditions. The timing manipulation was used to test the 319 effect of temporal proximity and grouping on the pre-activation of movements during 320 preparation. The relative compression and expansion of target IPIs by a scaling factor 321 of 2 in the fast and irregular timing conditions relative to the baseline condition (long 322 preparation duration and slow timing conditions) is in line with previous work on motor 323 timing (Wang et al., 2018). Although participants were trained to produce specific IPI 324 durations imposed by the target IPIs, relative timing, i.e., temporal IPI modulations 325 relative to the baseline condition, was key to evaluating the influence of timing at the 326 group and individual levels. Thus, relative timing was calculated offline from memoryguided Sequence trials (test phase) as each IPI duration (1st, 2nd, 3rd) relative to the 327 328 mean produced IPI duration in the baseline condition (in percent). Accordingly, relative 329 temporal error, was defined as the mean absolute deviation from the target IPI per trial 330 in percent.

In a *Sequence* trial, the *Go* cue remained on the screen for 3000 ms while in a *Probe* trial, the *Probe* cue for 1000 ms. This was followed by a fixation cross (1000 ms) and feedback (1000 ms) with a varying ITI of 500, 900 and 1300 ms. As a result, a *Sequence* trial was 6.5 min long and a *Probe* trial 4.5 min long. The participants underwent the same structure of training and test sessions as in Experiment 1. The timing conditions were equally matched to the number of all trial types in each block. Overall, in this experiment participants completed 2016 trials over 58 blocks.

Experiment 3 – Sequence timing and control movement. Procedures for Experiment 3 were identical to Experiment 1, except the introduction of additional *Probe* trials that cued the thumb. Thumb presses were not part of any target finger sequence. Thus, they served as a control condition to obtain reaction times and error rates for unplanned movements. Across each training stage, there were 60 *Probe* trials, while the test phase (30 blocks × 26 trials) contained 360 memory-guided *Sequence* trials (120 trials per timing condition), 360 *Probe* trials (30 trials per digit per

timing condition), and 60 thumb *Probe* trials (20 trials per timing condition). Overall,
participants completed 1990 trials over 72 blocks, excluding the practice block.

347 Feedback. In all experiments, a points system was designed to reward fast 348 initiation and accurate performance and avoid any performance drift in blocks with 349 motor production from memory. To incentivize the participants to gain as many points 350 as possible on each trial, we offered an extra monetary reward (£10) to those two with 351 the highest total points. In Sequence trials, points (0-10) could be awarded based on 352 three performance criteria: finger press accuracy, sequence initiation reaction time 353 (RT), i.e., response from Go cue to the first press, and temporal error (deviation from 354 the target IPIs). Points in each Sequence trial were the sum of the points for initiation 355 RT and mean temporal error, multiplied by finger press accuracy points (0 or 1). If at 356 least one incorrect press or an incorrect number of presses was recorded (< 4 or > 4), 357 0 points were given on that trial, regardless of initiation RT and temporal error. Points 358 gained from the initiation RT component of the sequence, were defined by tolerance 359 RT windows of 0-200, 200-360, 360-480, 480-560, 560-600 ms resulting in 5, 4, 3, 2 360 and 1 points, respectively. For late (> 600) responses, 0 points were given. Mean 361 temporal error was calculated for each trial as deviation of presses from target timing 362 in percent of the respective target IPI to account for the scalar variability of timing 363 (Jazayeri and Shadlen 2010; Rakitin et al. 1998). Thresholds for mean absolute 364 percentage deviation across all correct presses were set at 10, 20, 30, 40 and 50 % 365 assigning 5, 4, 3, 2 and 1 points, respectively. Mean temporal error above 50 % 366 resulted in 0 points.

Points (0-5) in each *Probe* trial were calculated based on finger press accuracy (0 or 1) and RT utilizing the same tolerance RT windows. In the case of an incorrect press or incorrect number of presses (< 1 or > 1), 0 points were given regardless of the RT length. The points were displayed on the screen after each *Probe* trial whilst after a *Sequence* trial they were presented above a schematic visual feedback.

372 Schematic feedback provided information on both finger press accuracy and 373 temporal error performance only after each *Sequence* trial. An '*x*' or a '-' symbol was 374 shown for every correct or incorrect press, respectively. For early presses, the 375 respective symbol was displayed below the midline (target timing), while for late 376 presses it was displayed above. For orientation, the lines above and below (upper and

lower border) corresponded to timing deviations as large as the target IPI itself (100%). 377 378 Timing deviation was only shown for second, third and fourth presses of the sequence. 379 The first symbol reflected the first press and was always positioned on the midline, 380 representing the starting point of the sequence. Participants were instructed to adjust 381 their performance by keeping the 'x' symbols as close to the midline as possible. 382 Deviation from the target onset (presented or assumed) rather than the interval timing 383 encouraged participants to synchronise with the instructed sequences during training. 384 however, may have contributed to a tendency to compress the overall sequence length 385 during the memory-guided Sequence trials.

386

Participant exclusion criteria

387 In each experiment, mean finger error rate (percent error trials out of total trials) 388 during sequence production from memory (memory-guided Sequence trials; test 389 phase) above three standard deviations of the group mean performance was 390 considered as outlier performance. This was to ensure that participants reached a 391 comparable skill level in sequence production. Additionally, it allowed for a sufficient 392 number of trials for RT analysis per participant, which included correct trials only. Data 393 exclusion was blind to the individual *Probe* trial performance and, thus, independent 394 of the measures analysed to test our hypotheses. In Experiment 1, the data of one 395 participant was excluded who showed 53.1% finger error in the short, 54.7% in the 396 intermediate and 53.9% in the long preparation duration conditions. Two participants' 397 data sets were removed from Experiment 2, one with 25% finger error in the slow timing and 18.8% in the irregular timing conditions, whilst the other showed 44.5% 398 399 finger error in the fast timing condition. The data of one participant was excluded from 400 Experiment 3 due to 12.5% finger error in the fast timing condition. Overall, the data 401 of 19 participants were analysed for Experiment 1, 18 participants for Experiment 2, 402 and 18 participants for Experiment 3.

403 Data analysis

Data analyses were performed using custom written code in MATLAB (v9.2 R2017a, The MathWorks, Inc., Natick, Massachusetts, United States), and SPSS v22.0 (IBM Corp., Armonk, N.Y., USA).

407 Sequence planning. Median reaction time (RT; correct trials only) and mean 408 error rate in *Probe* trials were used as dependent measures for assessing the 409 availability of movements corresponding to different sequence positions during 410 planning. First, we tested for the RT and error rate increases from 1st to 2nd, 2nd to 3rd 411 and 3rd to 4th positions in each experiment. These were tested in the baseline condition 412 common across the three experiments (long preparation duration and slow timing 413 conditions). One-tailed paired samples t-tests were performed on the raw RTs and 414 error rates, based on the one-sided hypothesis of an increase with position number. 415 The position-dependent differences for error were further examined in the lower and 416 upper RT quartiles to test for position-dependent increases of press error depending 417 on response speed.

418 Second, to test for the interaction of factors Position (1st, 2nd, 3rd, 4th) and 419 Preparation duration (short / 500 ms, intermediate / 1000 ms, long / 1500 ms) in Probe 420 trials of Experiment 1, the raw RTs and error rates were submitted to two-way repeated 421 measures ANOVAs. Using the same test, we assessed the interaction of the factors 422 Position and sequence Timing (slow, fast, irregular) in Experiments 2 and 3. Significant 423 interaction effects were investigated using planned repeated contrasts, to determine 424 the changes relative to baseline that were driving the interaction. To evaluate the RT 425 and error rate for the control movement (Experiment 3), we used two-tailed paired 426 samples t-tests (control vs 4th position).

427 Third, we calculated the increase of RT and error rate for each probed position 428 relative to the first position in each condition (in %) for each participant. This enabled 429 us to quantify and visualise the relative position-dependent increases in each condition (Figure 2). Further, we calculated the average relative RT and error differences 430 between adjacent positions (mean difference across 1st minus 2nd, 2nd minus 3rd, 3rd 431 432 minus 4th) in the baseline condition for each participant as markers of the movements' 433 pre-activation gradient size during sequence planning. One-way repeated measures 434 ANOVAs in each experiment were used to assess modulations of the latter by the

experimental conditions (Preparation duration in Experiment 1 and Timing in 435 436 Experiments 1 and 2). To test for the association between these measures and 437 sequence performance (initiation RT of correct sequences, relative temporal error, and 438 finger error rate) six one-tailed Pearson's correlation analyses were performed across 439 experiments (N = 55). Further, a median split was calculated based on each 440 performance measure for raw mean RTs and error rates for each position in the 441 baseline condition. These were subjected to three mixed ANOVAs (Position × Group) 442 to test for the position-dependent differences in movement availability during planning 443 depending on performance (N = 55).

Finally, we looked at the percent of presses associated with the 1st, 2nd, 3rd, and 445 4th positions of the planned sequence in erroneous *Probe* trials, for each probed 446 position separately (four one-way repeated measures ANOVAs; N = 55).

447 Sequence production. Only the memory-guided Sequence trials (test phase) 448 were used for analysing the components of sequence production. First, relative timing (percent duration of each IPI relative to the mean produced IPI in the baseline 449 450 condition) was subjected to a 3 x 3 repeated measures ANOVA, for each experiment, depending on IPI (1st, 2nd, 3rd) and Preparation duration (Experiment 1) or sequence 451 452 Timing (Experiments 2 and 3). Finally, to evaluate the fluency and accuracy of 453 sequence production, we calculated sequence initiation RT (online recording of Go 454 cue to first press latency), relative temporal error (deviation from target IPI) and finger 455 press error (percent trials with incorrect presses). These constituted the three 456 performance measures to reflect skill in sequence execution and were analysed for 457 each experiment separately in nine one-way repeated measures ANOVAs to assess 458 modulations of skill by Preparation duration or Timing.

The error data of both *Probe* and *Sequence* trials were arcsine transformed (Winer et al. 1991) before they were submitted to the ANOVA models and t-tests due to violation of normality. Partial eta squared ratios and Cohen's *d* are reported as measures of effect sizes in the corresponding tests.

463 **Results**

464 Availability of movements during sequence planning is dependent on 465 their position in the planned sequence

In all three experiments, participants were trained for two days to associate 466 467 abstract visual cues with four-element finger sequences. They were instructed to 468 produce the sequences with a particular temporal structure (Timing: slow, fast, 469 irregular) following a brief preparation period (Preparation duration: short / 500ms, 470 intermediate / 1000 ms, long / 1500 ms). In half of the trials in the test phase (Day 3), 471 a *Probe* cue instructed participants to respond with the corresponding finger press as 472 quickly and accurately as possible at the end of the planning phase (Figure 1c). This 473 allowed us to probe the availability of movement associated with each position of the 474 planned sequence (1st - 4th) for accurate and fast execution. Based on our previous 475 neurophysiological findings (Kornysheva et al. 2019) in a similar task that showed a 476 graded pre-activation of movements during planning according to their sequential 477 position, we hypothesized that the behavioural availability of movements during 478 planning will be position-dependent. Specifically, we predicted a significant increase in RT and error rate for probed movements from 1st to 2nd and 2nd to 3rd positions. 479 480 Based on our neurophysiological results we did not expect an increase in movement availability from penultimate to final position (here: 3rd to 4th), but the latter has been 481 482 previously observed in the context of a drawing sequence task in non-human primates 483 (Averbeck et al. 2002). Additionally, we included probes for a control movement 484 (Experiment 3) to reveal whether the movement associated with the last position of 485 the planned sequence is more accurately and quickly selected and executed than a 486 movement that is not part of the sequence. A higher behavioural availability of the last 487 position movement would suggest that the sequence movements are more pre-488 activated, albeit to a different level, rather than activated and inhibited relative to a 489 baseline movement. Position-dependent RT and press error increases were analysed 490 from trials in the experimental condition which constituted the baseline in all three 491 experiments (long preparation duration - 1500 ms - and slow timings).

492 *Reaction times to movement probes.* Figure 2a shows the percent RT increase
493 relative to the RT for the movements associated with the first position, respectively (cf.
494 Supplemental Figure S2a for raw RT values; Supplemental Table S1a for statistics).

Experiment 1 revealed a significant RT increase from 1st to 2nd position (paired 495 samples t-test: t (18) = -7.45, p < .001, d = 1.32, one-tailed) but not from 2nd to 3rd 496 position (t (18) = .05, p = .479, d = .01) or from 3rd to 4th position (t (18) = -.72, p = .241, 497 498 d = .09). Experiment 2 replicated the RT results from Experiment 1 revealing a significant RT increase from 1st to 2nd position (t (17) = -6.45, p < .001, d = 1.60), but 499 not from 2^{nd} to 3^{rd} (t (17) = -.63, p = .267, d = .16) or 3^{rd} to 4^{th} position (t (17) = -.25, p 500 = .404, d = .05). Experiment 3 showed a significant RT increase from 1st to 2nd position 501 (t (17) = -4.61, p < .001, d = 1.03) and, unlike the Experiments 1 and 2, also from 2nd 502 503 to 3^{rd} position (t (17) = -2.41, p = .014, d = .40). As in Experiments 1 and 2, the RT 504 increase from 3^{rd} to 4^{th} position was not significant (t(17) = -.21, p = .417, d = .04). To further investigate whether the inconsistent mean RT increase for probes from 2nd to 505 506 3rd position would be resolved with higher power, a pooled analysis across the three 507 experiments was performed (N = 55). This revealed a marginal RT increase from 2^{nd} 508 to 3^{rd} position (t (54) = -1.55, p = .063, d = .15), suggesting that this overall increase 509 was highly variable across subjects. Finally, the RT of the control movement was 510 significantly higher than the movement associated with the last position (4th) of the 511 planned sequence (paired samples t-test: t(17) = 3.04, p = .007, d = .86, two-tailed).

512 Across experiments, the present RT data shows that during sequence planning, 513 correct finger presses associated with earlier positions in a sequence can be selected 514 and executed quicker than those associated with later positions, suggesting a position-515 dependent pre-activation gradient. In particular, the latter can switch flexibly trial-by-516 trial, depending on which finger sequence is retrieved and planned in a particular trial. 517 The data also suggests that the availability is modulated up to three positions ahead. 518 with RT increases for later positions becoming less consistent across subjects. Finally, 519 although the movement associated with the last position was the slowest to execute 520 on average, it was still faster than a control movement not featuring in the planned 521 sequence.

Error rates to movement probes. Figure 2b shows the percent press error increase relative to the error rates for the movements associated with the first position, respectively (cf. Supplemental Figure S2b for raw press error rates; Supplemental Table S1a for statistics). Experiment 1 revealed significant error increases from 1st to 2nd position (paired samples t-test: *t* (18) = -6.65, *p* <.001, *d* = 1.83, one-tailed) and

from 2^{nd} to 3^{rd} position (*t* (18) = -1.93, *p* = .035, *d* = .27), and no significant increase 527 from 3^{rd} to 4^{th} position (t (18) = -1.24, p = .116, d = .21). Experiment 2 replicated the 528 significant error increase from 1st to 2nd cf. position (t(17) = -5.51, p < .001, d = 1.57) 529 and from 2^{nd} to 3^{rd} position (t(17) = -2.05, p = .029, d = .43). In contrast, the difference 530 from 3rd to 4th position showed no significant increase, but an unexpected decrease of 531 errors (t (17) = 2.60, p = .010, d = .54). Experiment 3 again replicated the significant 532 error increases from 1st to 2nd position (t(17) = -7.77, p < .001, d = 1.83) and from 2nd 533 to 3^{rd} position (t (17) = -1.88, p = .039, d = .58), whilst there was no significant 534 difference between the 3rd and 4th positions (t(17) = .77, p = .227, d = .20). The control 535 movement did not show a significant increase in errors compared to the 4th position 536 537 (paired samples t-test: t(17) = -.81, p = .430, d = .26, two-tailed).

538 The error rate data from all experiments indicate that during sequence planning, 539 movement probes associated with earlier positions in a sequence are more likely to 540 lead to correct finger presses than those associated with later positions, which are 541 more prone to erroneous finger presses. Like RT, error rate data points to a position-542 dependent pre-activation gradient for movements associated with the first three 543 positions in the sequence, but respective error increases between the first 3 positions 544 appear to be more pronounced and consistent across participants, particularly for increases from 2nd to 3rd position. Further, it shows that movements associated with 545 the last (4th) position are equally error prone as a sequence irrelevant control 546 547 movement, although the former is still faster to execute when selected correctly. Taken 548 together, our findings advocate the presence of a preparatory pre-activation gradient 549 which renders movements associated with later sequence positions less available for 550 correct selection and fast execution. They point to the planning of up to three constituent movements in advance within a brief preparation period and retrieval from 551 552 memory. This pre-activation level does not increase linearly with movement positions but falls off and becomes more variable across participants for movements associated 553 554 with later positions. The variability of the gradient during planning across participants 555 is examined below in the context of skilled performance.

556Position-dependent differences in movement availability are modulated557by preparation duration, not timing

558 Next, we examined whether the position-dependent availability for correct 559 movement selection and fast execution during planning is modulated by the time to 560 prepare a sequence, or the planned sequence timing.

561 Preparation duration. According to our accuracy hypothesis, a more accurate plan of the sequence progressively established across preparation durations of 500-562 563 1500 ms would lead to an expansion of the pre-activation gradient (Kornysheva et al. 564 2019). In Experiment 1 (cf. Supplemental Table S1b for statistics), we found a large significant interaction of Position and Preparation duration for error rates (4 × 3 565 repeated measures ANOVA of raw press error rates: F (6, 108) = 3.35, p = .005. 566 567 np^2 = .16). The latter was driven by a significant error rate increase for 2nd relative to 568 1st sequence positions with longer preparation duration (500 *vs* 1500 ms preparation 569 duration, F(1, 18) = 15.89, p = .001, $\eta p^2 = .47$). This contrast was also significant for 570 RTs (F(1, 18) = 5.89, p = .026, $\eta p^2 = .25$), although the interaction between Position 571 and Preparation duration for RTs did not reach significance (4 × 3 repeated measures ANOVA of raw RTs: F(6, 108) = 2.07, p = .063, $\eta p^2 = .10$). This shows that the increase 572 in RT and error rate from 1st to 2nd position became more pronounced with longer 573 574 preparation durations, an effect which drove the significant interaction.

575 Importantly, both the relative RT and error differences became more 576 pronounced with longer preparation duration conditions (one-way repeated measures 577 ANOVA of: Relative RT differences - Experiment 1, F(2, 36) = 4.38, p = .020, $\eta p^2 = .20$; Relative error differences - Experiment 1, F(2, 36) = 3.46, p = .042, 578 np^2 = .16; cf. Supplemental Table S1c for statistics). Thus, more time to prepare the 579 580 sequence made the probed movements associated with later positions less available 581 for correct selection and fast execution, and vice versa. This suggests that the pre-582 activation state of the planned movements became more differentiated according to 583 position and the pre-activation gradient expanded across the sequence retrieval and 584 preparation period.

585 *Timing*. According to the timing hypothesis, movements in a sequence that are 586 closer in time should have more similar levels of pre-activation, and vice versa, leading 587 to a contraction and expansion of the pre-activation gradient for each action. Contrary 588 to the timing hypothesis, the interaction between Position and Timing (cf. 589 Supplemental Table S1b for statistics) did not reach significance, neither for RTs, nor

for error rate increases (4 × 3 repeated measures ANOVA of: Raw RTs - Experiment 590 2, F (3.27, 55.54) = 2.30, p = .082, $\eta p^2 = .12$, Greenhouse-Geisser corrected, χ^2 (20) 591 = 42.61, p = .003; Experiment 3, F (3.87, 65.79) = .98, p = .426, ηp^2 = .05, Greenhouse-592 593 Geisser corrected, $\chi^2(20) = 34.06$, p = .028; Raw error rates - Experiment 2, F (6, 102) = 1.86, p = .095, $\eta p^2 = .10$; Experiment 3, F (6, 102) = 1.02, p = .416, $\eta p^2 = .06$). This 594 finding was corroborated by an absent effect of Timing on either the relative RT or the 595 596 relative error differences (one-way repeated measures ANOVA of: Relative RT 597 differences - Experiment 2, $F(1.48, 25.23) = .68, p = .475, np^2 = .04$, Greenhouse-598 Geisser corrected, x^2 (2) = 6.83, p = .033; Experiment 3, F (2, 34) = 1.92, p = .162. 599 $np^2 = .10$; Relative error differences - Experiment 2, F (2, 34) = .00, p = .999, $np^2 = .00$; Experiment 3, F(1.27, 21.52) = 1.50, p = .241, $\eta p^2 = .08$, Greenhouse-Geisser 600 601 corrected, χ^2 (2) = 13.87, p = .001; cf. Supplemental Table S1c for statistics). We 602 investigated whether the results may be contaminated by participants that 603 considerably deviated in their relative temporal error performance (memory-guided Sequence trials; test phase). Therefore, we performed the same analyses after 604 605 removing outlier participants that showed little modulation of timing during sequence 606 production (cf. Supplemental Figure S3). However, without these outliers, the 607 interaction between Position and Timing was still not significant. Overall, these 608 analyses indicate that preparing a sequence that is twice as fast, or temporally 609 grouped, did not impact the position-dependent pre-activation gradient of movements 610 during sequence planning.

611 **Position-dependent modulation of press error during planning is revealed** 612 **through fast responses to probes**

613 Next, we sought to determine whether the characteristic position-dependent 614 increases in press errors in *Probe* trials were driven by automatic responses to *Probe* 615 cues, or by deliberate movement selection. To investigate this guestion, we analysed the position-dependent error increases for the first versus last RT distribution guartiles 616 617 in each participant (baseline condition: 1500 ms preparation duration and slow timing). 618 Figure 2c (cf. Supplemental Table S1a for statistics) illustrates the press error 619 increases relative to the first position for fast and slow RT quartiles. In fast response 620 trials, we found significant error increases up to the 3rd position in Experiments 1 and 3 (paired samples t-tests: Experiment 1, 1st to 2nd position, t (18) = -6.54, p <.001, d 621 622 = .54, one-tailed; 2^{nd} to 3^{rd} position, t (18) = -2.87, p = .005, d = .40; 3^{rd} to 4^{th} position, t(18) = 3.12, p = .003, d = .48; Experiment 3, 1st to 2nd position, t(17) = -6.59, p < .001,623 d = 2.12; 2nd to 3rd position, t(17) = -1.82, p = .043, d = .55; 3rd to 4th position, t(17) = -1.82, p = .043, d = .55; 3rd to 4th position, t(17) = -1.82, p = .043, d = .55; 3rd to 4th position, t(17) = -1.82, p = .043, d = .55; 3rd to 4th position, t(17) = -1.82, p = .043, d = .55; 3rd to 4th position, t(17) = -1.82, p = .043, d = .55; 3rd to 4th position, t(17) = -1.82, p = .043, d = .55; 3rd to 4th position, t(17) = -1.82, p = .043, d = .55; 3rd to 4th position, t(17) = -1.82, p = .043, d = .55; 3rd to 4th position, t(17) = -1.82, p = .043, d = .55; 3rd to 4th position, t(17) = -1.82, p = .043, d = .55; 3rd to 4th position, t(17) = .55; 3rd position, t(17) = .55; 624 1.63, p = .061, d = .35) and up to the 2nd position in Experiment 2 (1st to 2nd position, t 625 $(17) = -6.99, p < .001, d = 1.57; 2^{nd}$ to 3^{rd} position, $t(17) = -.93, p = .184, d = .43; 3^{rd}$ to 626 627 4^{th} position, t(17) = 1.43, p = .085, d = .54). In contrast, in slow response trials, errors 628 did not increase significantly with position (Experiment 1, 1^{st} to 2^{nd} position, t(18) = .59, p = .281, d = .20; 2nd to 3rd position, t (18) = -.55, p = .294, d = .19; 3rd to 4th position, 629 t (18) = -.60, p = .277, d = .16; Experiment 2, 1st to 2nd position, t (17) = -.57, p = .290, 630 d = .20; 2nd to 3rd position, (t(17) = .00, p = .500, d = .00; 3rd to 4th position, t(17) = .57, 631 p = .290, d = .20; Experiment 3, 1st to 2nd position, t(17) = -.34, p = .368, d = .08; 2nd632 to 3^{rd} position, t(17) = .15, p = .443, d = .04; 3^{rd} to 4^{th} position, t(17) = .54, p = .299, 633 634 d = .17). The control movement did not show more errors than the 4th position in either 635 fast or slow RT as in the main results (Fast RTs, t(17) = -.95, p = .353, d = .28, two-636 tailed; Slow RTs, t(17) = .10, p = .922, d = .03).

637 These results demonstrate that the position-dependent availability of 638 movements for correct selection following movement *Probe* cues is driven by 639 automatic responses rather than by a cognitive selection process.

640

641

Incorrect presses to movement probes during planning are dominated by the movement in the first sequence position

642 We investigated whether incorrect presses in Probe trials were associated with specific positions of the planned sequence on that trial (Figure 3; cf. Supplemental 643 644 Table S2 for statistics). This was undertaken for each probed position separately and across all three experiments. Results for 1st position (Figure 3, upper left) did not yield 645 significant differences among the press rate for 2nd, 3rd, and 4th positions (one-way 646 647 repeated measures ANOVA: F(2, 108) = .63, p = .535, $np^2 = .01$). In contrast, probing 648 the movements associated with 2nd, 3rd, and 4th positions revealed that participants 649 consistently selected the 1st position more frequently. Specifically, when the 2nd 650 position was probed (Figure 3, upper right), there was a significant difference among 651 1st, 3rd, and 4th erroneously pressed positions (F (1.38, 74.36) = 84.70, p < .001, 652 ηp^2 = .61, Greenhouse-Geisser corrected, χ^2 (2) = 31.92, p < .001; 1st position higher than 3^{rd} position, p < .001; 1^{st} position higher than 4^{th} position, p < .001; 3^{rd} position 653 higher than 4th position, p = .007). Similarly, the press rate for the 1st position when the 654 3rd position was probed (Figure 3, lower left) was higher than the 2nd and 4th pressed 655 656 positions (F(1.34, 72.50) = 84.90, p < .001, $np^2 = .61$, Greenhouse-Geisser corrected, x^2 (2) = 35.65, p < .001; 1st position higher than 2nd position, p < .001; 1st position 657 higher than 4th position, p < .001; 2nd position marginally lower than 4th position, p658 659 = .069). The 4th probed position (Figure 3, lower right) produced higher 1st position presses (F (1.54, 83.34) = 42.95, p < .001, $\eta p^2 = .44$, Greenhouse-Geisser corrected, 660 χ^2 (2) = 18.60, p < .001; 1st position higher than 2nd position, p < .001; 1st position 661 higher than 3^{rd} position, p < .001; 2^{nd} position not significantly higher than 3^{rd} position, 662 663 p = 1.000).

664 The distribution of erroneous presses shows that the movement availability was 665 highly biased towards the production of the movement in the first position in each 666 respective sequence upon retrieval and planning of the cued sequence.

667

Greater position-dependent differences in movement availability during 668 planning predict better performance

Position-dependent pre-activation differences between sequential movement 669 670 patterns during planning have been shown to predict the participants' subsequent 671 performance accuracy (Kornysheva et al. 2019). Specifically, the distance (i.e.,

difference) between the neural pattern probabilities of consecutive movements during 672 planning predicted more skilled sequence execution. Accordingly, we predicted that 673 674 larger position-dependent differences in availability of movements for correct selection 675 and fast execution during planning would correlate with a more skilled performance 676 during sequence execution. Position-dependent differences in availability of movements was considered a proxy measure for the pre-activation gradient size (cf. 677 678 relative RT and error differences in Data analysis, Methods). We took faster initiation 679 of correct sequences after the Go cue, as well as reduced relative temporal errors and finger errors as markers of a more skilled performance. Correlation analyses were 680 681 performed on group data (N = 55) obtained from trials in the baseline condition present 682 in all experiments (long preparation duration and slow timing conditions; Figure 4a, b; 683 cf. Supplemental Figure S4 for raw RT and error differences; Supplemental Table S3a 684 for statistics). Results showed that participants with larger relative RT and error 685 differences during planning initiated correct sequences faster (Relative RT differences: 686 r = -.39, p = .002; Relative error differences: r = -.54, p < .001, one-tailed). Larger 687 relative RT differences during planning were also correlated with lower relative 688 temporal error (r = -.35, p = .005). This association did not hold up for the relative error 689 differences (r = -.05, p = .356). Thus, the latter may be a less sensitive predictor for 690 temporal accuracy than the relative RT differences. In contrast to our predictions, we 691 did not find an association with finger error (Relative RT differences: r = .08, p = .273; 692 Relative error differences: r = .12, p = .196). This was likely due to ceiling effects in finger press accuracy performance attributable to the limited number of trained finger 693 694 sequences.

695 To inspect the position-dependent slopes in movement availability based on 696 sequence performance, we performed median split-based initiation RT, relative 697 temporal error, and finger error (Figure 4 insets; cf. Supplemental Table S3b for 698 statistics). Participants with faster initiation RTs exhibited larger position-dependent 699 RT differences (Figure 4a, inset) compared to those with slower initiation RTs (mixed ANOVA with median split of initiation RT: Main effect of Group, F(1, 53) = 33.63, p 700 < .001, $\eta p^2 = .39$; Position × Group, F (3, 159) = 5.70, p = .001, $\eta p^2 = .10$). Equally, the 701 position-dependent press error differences (Figure 4b, inset) were steeper for 702 703 participants with fast initiation RTs (mixed ANOVA with median split of initiation RT:

Main effect of Group, F(1, 53) = 10.77, p = .002, $\eta p^2 = .17$; Position × Group, F(3, 159) = 3.90, p = .010, $\eta p^2 = .07$). Median splits by relative temporal error or finger error did not show differences in movement availability during planning, confirming further that this relationship is either more subtle (temporal error) or absent (finger error).

Together, these analyses show that behavioural markers of a more expanded pre-activation gradient can predict faster initiation of correct finger sequences and improved relative temporal, but not finger accuracy during production.

Next, we conducted a series of extended analyses focussing on sequence production. These additional analyses examined whether participants – on average – produced the sequences from memory with accurate relative timing, and whether preparation time and sequence timing conditions changed performance, i.e., speed of correct sequence initiation, as well as temporal and finger accuracy.

716 *Participants produced sequences from memory with correct relative* 717 *timing*

Participants were trained to either retain the same (Experiment 1) or consistently modulate (Experiments 2 and 3) the relative timing during sequence production across sequence conditions. On average, participants produced the sequences with timing relative to the target inter-press-intervals (IPI) (Figure 5a; cf. Supplemental Table S4 for statistics; Supplemental Figure S5 for mean absolute press timing per trial).

724 The mean relative timing of finger presses in Experiment 1 was nearly identical across preparation duration conditions (Figure 5a, left). Nevertheless, we detected a 725 726 small but significant interaction between IPI and Preparation duration (3 × 3 repeated measures ANOVA: F(4, 72) = 2.53, p = .048, $\eta p^2 = .12$), explained by IPI modulations 727 728 of 9 ms across conditions. Post-hoc comparisons (Bonferroni-corrected for nine tests) 729 revealed a significant shortening of the 1st interval in the short preparation duration (p = .002) and of the 1st (p = .002) and 3rd (p = .004) intervals in the intermediate 730 731 compared to the long preparation duration. This shows that there was a tendency to slightly compress the 1st and 3rd intervals with shorter preparation time. If there were 732 733 a timing confound on sequence planning duration in Experiment 1, the timing effect 734 should have been vastly amplified by the experimental modulation of timing requiring

the doubling or halving of IPIs in Experiments 2 and 3. However, we did not observeany strong and consistent effect of the latter on sequence planning.

737 Experiment 2 (Figure 5a, middle) showed a large significant interaction of IPI and Timing $(3 \times 3 \text{ repeated measures ANOVA: } F(1.26, 21.42) = 59.49, p < .001,$ 738 np^2 = .78, Greenhouse-Geisser corrected, $\chi^2(9) = 97.83$, p < .001), in line with the task 739 740 instructions. The pairwise comparisons (Bonferroni-corrected for nine tests) of the 741 produced IPIs confirmed that the participants modulated their relative timing according 742 to the target IPI structure. In accordance with the cued sequence, the 1st IPI was 743 significantly longer in the slow than in the fast (p < .001) and the irregular timing 744 conditions (p < .001), while it did not differ in the fast vs irregular timing conditions (p = .001) 1.000). The 2nd IPI length increased slightly, yet proportionally for both the slow and 745 746 fast timing conditions, retaining the significant difference (p < .001), and doubled in 747 length in the irregular relative to the slow timing condition (p < .001). The 3rd IPI 748 exhibited a very similar profile to the 1st IPI (slow vs fast, p < .001; slow vs irregular, 749 p < .001), but its length decreased slightly in the fast compared to the irregular timing 750 condition (p = .027). Experiment 3 (Figure 5a, right) replicated the findings of 751 Experiment 2 showing a significant interaction of IPI and Timing (3 × 3 repeated 752 measures ANOVA: F (1.56, 26.49) = 17.37, p < .001, $np^2 = .51$, Greenhouse-Geisser 753 corrected, $\chi^2(9) = 61.31$, p < .001). Again, post-hoc pairwise comparisons (Bonferroni-754 corrected for nine tests) confirmed that the 1st IPI in the slow timing was longer than 755 that in the fast (p = .001) and irregular (p = .003) timing conditions, while no difference was found between the fast and irregular timing conditions (p = 1.000). The 2nd IPI was 756 757 significantly longer in the slow compared to the fast timing condition (p = .001), but shorter compared to the irregular timing condition (p = .005). Similarly, the 2nd IPI in 758 759 the fast timing was half as long than in the irregular timing condition (p < .001). The 3rd 760 IPI was twice as long in the slow compared to the fast timing condition (p < .001). It 761 did not show a significant shortening for the irregular timing when compared to the 762 slow timing condition (p = 1.000) and showed only a marginally significant difference 763 between the fast and irregular timing conditions (p = .096).

764 Overall, these results demonstrate that, on average, participants produced the 765 finger sequences from memory with accurate relative timing across conditions.

766

Longer preparation durations shortened initiation of correct sequences

We found a significant difference in sequence initiation RT with Preparation 767 duration (one-way repeated measures ANOVA: Experiment 1, F (1.38, 24.88) = 768 52.81, p < .001, $np^2 = .75$, Greenhouse-Geisser corrected, χ^2 (2) = 10.07, p = .006) 769 770 (Figure 5b, left; cf. Supplemental Table S4 for statistics). Pairwise comparisons 771 (Bonferroni-corrected for three tests) confirmed that sequence initiation RT was 772 significantly faster for the intermediate (1000 ms) and long (1500 ms) preparation 773 duration than following a short (500 ms) preparation duration (intermediate vs short, p 774 < .001; long vs short, p < .001). Further, sequence initiation RT following a long 775 preparation duration was significantly faster as compared to the intermediate 776 preparation duration (p = .005). In experiments with single movements the effect of 777 variable preparation duration on RT is known as the foreperiod effect (Foley 1959; 778 Vallesi et al. 2007). It can be accounted for by generic motor preparedness due to 779 heightened temporal expectation (hazard rate) for longer preparation durations (Bueti 780 et al. 2010), and includes carry-over effects across trials (Langner et al. 2018; 781 Steinborn and Langner 2012) (cf. Supplemental Figure S6 for preparation duration 782 effects of preceding trials in Experiment 1). However, the effect on initiation RT 783 reported here cannot be attributed to general temporal preparedness alone. In contrast 784 to classical foreperiod paradigms the current paradigm involves a Sequence cue at 785 the start of the foreperiod, instead of a neutral warning signal. Therefore, a facilitation 786 of initiation RT will reflect the state of sequence preparedness that increases with 787 longer durations (Ariani and Diedrichsen 2019; Sternberg et al. 1978), not just non-788 specific effects of temporal expectation.

789 There was no main effect of Timing on sequence initiation RT in Experiment 2 (one-way repeated measures ANOVA: $F(1.41, 23.92) = 1.70, p = .207, np^2 = .09,$ 790 Greenhouse-Geisser corrected, χ^2 (2) = 8.76, *p* = .013), but a main effect of Timing in 791 792 Experiment 3 (one-way repeated measures ANOVA: F(1.29, 21.99) = 11.59, p = .001, 793 np^2 = .41, Greenhouse-Geisser corrected, χ^2 (2) = 12.63, p = .002). As explained by 794 pairwise comparisons (Bonferroni-corrected for three tests), participants in Experiment 3 were slower at initiating a sequence of slow timing when compared to fast timing (p 795 796 = .006) and irregular timing (p = .010). There was no difference in initiation RT between 797 the fast and the irregular timing conditions (p = .118). This effect was not consistent 798 across Experiments 2 and 3, but present at the mean level in both experiments. This implies that sequences with a slow isochronous timing structure were less prepared
for initiation following a *Go* cue compared to sequences that started with two presses
in short succession (fast and irregular timing structures), which may be more prone to
a rushed initiation.

803 Sequences involving irregular inter-press-intervals were produced with 804 less accurate timing

805 Next, we established whether preparation duration (Experiment 1) and 806 sequence timing (Experiments 2 and 3) modulated relative temporal error during 807 sequence production (Figure 5b, middle: cf, Supplemental Table S4 for statistics). In 808 Experiment 1, mean relative temporal error did not differ among the three preparation 809 duration conditions (one-way repeated measures ANOVA: F(2, 36) = .11, p = .901,810 $\eta p^2 = .01$). Here relative temporal performance may have been compensated in the 811 short preparation duration condition by slower initiation RT (cf. above). In Experiment 812 2, there was a significant effect of Timing (one-way repeated measures ANOVA: F (2, 813 34) = 28.23, p < .001, $np^2 = .62$). Pairwise comparisons (Bonferroni-corrected for three 814 tests) revealed that participants performed at a lower relative temporal error when 815 producing a sequence of slow timing compared to irregular timing (p < .001) and a 816 sequence of fast timing compared to irregular timing (p < .001), while there was no 817 difference between sequences in the slow vs fast timing conditions (p = 1.000). 818 Experiment 3 replicated the main effect of Timing (one-way repeated measures ANOVA: $F(1.45, 24.72) = 7.06, p = .007, \eta p^2 = .29$, Greenhouse-Geisser corrected, 819 820 $\chi^2(2) = 7.53$, p = .023). In line with the findings of Experiment 2, there were less relative 821 temporal errors in the slow timing (p = .049) and fast timing (p = .008) conditions when 822 compared to the irregular timing condition. Again, there was no significant difference 823 in relative temporal performance between the two isochronous conditions (slow vs 824 fast, p = 1.000). In sum, the production of sequences which consisted of nonisochronous IPIs (irregular timing condition) as opposed to equal IPI lengths 825 826 (isochronous timing conditions; slow, fast) were associated with decreased relative 827 temporal accuracy.

828 Finger press accuracy in sequences produced from memory was 829 matched across conditions

830 In the test phase, participants produced finger press sequences entirely from 831 memory. Neither Preparation duration (one-way repeated measures ANOVA: 832 Experiment 1, F(2, 36) = .23, p = .795, $\eta p^2 = .01$), nor Timing (one-way repeated 833 measures ANOVA: Experiment 2, F(2, 34) = .02, p = .984, $np^2 = .00$; Experiment 3, 834 F(2, 34) = .96, p = .394, $\eta p^2 = .05$) affected finger error during sequence production 835 (Figure 5b, right; cf. Supplemental Table S4 for statistics). This means that participants 836 prepared the finger order of cued sequences with the same accuracy, regardless of 837 the preparation time or temporal structure of the planned sequence. Note that finger 838 error in sequence production was higher in Experiment 1 than in Experiments 2 and 839 3. This is likely due to Experiment 1 involving sequences of two different finger 840 sequences on a trial-by-trial basis, whereas Experiments 2 and 3 involved the same 841 finger sequence performed with different timing.

842

843 Discussion

844 Sequence planning is central to skilled action control, however its content and 845 structure is poorly understood (Bullock 2004; Remington et al. 2018). 846 Neurophysiological findings have demonstrated that a trained movement sequence is 847 pre-planned by establishing a competitive pre-activation gradient of movement 848 patterns according to their serial position, and that the guality of this neural pattern 849 during planning predicts subsequent performance (Averbeck et al. 2002; Basu and 850 Murthy 2020; Kornysheva et al. 2019; Pinet et al. 2019). Here we report a putative 851 behavioural marker of this competitive pre-activation gradient. During a short retrieval and preparation period, we measured the behavioural availability of each constituent 852 853 movement of the planned sequence for accurate and fast production. Our findings 854 show that behavioural availability is dependent on the sequence position the 855 respective movements are associated with, mirroring the pre-activation gradient 856 observed in neurophysiological studies (Averbeck et al. 2002; Kornysheva et al. 2019) 857 as predicted by competitive queuing (CQ) models (Bullock 2004; Burgess and Hitch 858 1999; Hartley et al. 2016; Hartley and Houghton 1996). Critically, a stronger 859 differentiation between the state of movements assigned to consecutive sequence 860 positions correlated with markers of skilled production – the speed of correct sequence 861 initiation and the temporal production accuracy. In contrast, the latter did not reliably reflect the sequence production speed, or the inter-press-interval pattern of the planned sequence.

864 Sequence planning markedly contrasts with mechanisms for non-sequential 865 movement planning involving multiple movement options: In the latter, a cued set of 866 possible movements triggers equal activity increase in cortical populations tuned to 867 the respective movements, and the preparatory competition is only resolved once a 868 cue specifies the target movement (Cisek and Kalaska 2005). In contrast, sequence 869 planning established a fine-tuned gradient of movement pre-activations, with the latter 870 switching flexibly on a trial-by-trial basis, in line with the retrieved sequence. Notably, 871 movements that were part of the planned sequence were executed faster than a 872 control movement which was not part of the retrieved sequence (Figure 2a, right). This 873 suggests that all constituent movements were concurrently pre-activated above a 874 passive baseline, albeit to a different degree depending on their position in the planned 875 sequence.

876 Our study provides a measure of the competitive state of constituent 877 movements *prior to* sequence production. This is complementary to previous 878 behavioural work which supports the presence of competitive queuing of sequence 879 presses *during* production, such as accuracy and RT curves obtained from sequence 880 execution (Rhodes et al. 2004; Verwey and Abrahamse 2012), or on-the-fly movement 881 planning following sequence initiation, assessed behaviourally (Behmer and Crump 882 2017) and through measures of cortico-spinal excitability (Behmer et al. 2018). Gilbert 883 and colleagues have employed a paradigm at the interface between sequence 884 preparation and production to characterize the competitive queuing profile of the 885 respective sequential movements - silent rehearsal (Gilbert et al. 2017). Here 886 participants were asked to listen to sequences of spoken digits and silently rehearse 887 the items during a retention interval. They received explicit instructions to rehearse the 888 sequence at the same pace as active production. After an unpredictable delay, a tone 889 prompted the report of an item being rehearsed at that moment and revealed graded 890 overlapping probabilities of neighbouring items, suggesting potential CQ during 891 internal rehearsal. In contrast to the latter study, our paradigm did not enable active 892 rehearsal during preparation: First, our participants retrieved the sequence entirely 893 from memory without a sensory instruction period which might have facilitated active

894 entrainment with the sequence prior to planning. Second, the period for sequence 895 retrieval and planning was comparatively brief (ranging from 500 to 1500 ms after 896 Sequence cue onset) and not sufficient to cycle through the full sequence at the rate 897 participants employed for active production. In addition, if the observed CQ gradient 898 were somehow driven by silent rehearsal at the target rate, it would have been more 899 pronounced for the fast sequences, as more of the planned sequence could fit into the 900 preparation phase. However, there was no significant difference between relative 901 availability of probed movements for fast and slow sequences.

902 Whilst active motor rehearsal at scale during the short preparation phase is 903 unlikely, an alternative serial preparation mechanism may be related to rapid sequence 904 replay. The latter has been observed in the hippocampus during navigation tasks 905 (Ólafsdóttir et al. 2018) and perceptual sequence encoding (Liu et al. 2019), as well 906 as in the motor cortex in the context of motor sequence learning tasks (Eichenlaub et 907 al. 2020). Replay has been shown to involve fast sweeps through the neural patterns 908 associated with the sequence during wakeful rest and planning (preplay) (Dragoi and 909 Tonegawa 2011; Drieu and Zugaro 2019; Jafarpour et al. 2014; Ólafsdóttir et al. 2018), 910 and is characterized by a multifold temporal compression (Eichenlaub et al. 2020; 911 Kurth-Nelson et al. 2016; Liu et al. 2019; Michelmann et al. 2019). How replay could 912 translate into a parallel pre-activation of serial movements reported here is uncertain. 913 One possibility is that serial sweeps during motor sequence planning involve fast 914 repeated replay fragments (Davidson et al. 2009; Michelmann et al. 2019) of different length during preparation, starting with the first elements – e.g. 1st-2nd-3rd, 1st-2nd, 1st, 915 1st-2nd-3rd-4th, 1st-2nd etc. This would produce an overall bias towards the pre-activation 916 917 of earlier rather than later parts of the planned sequence. This, in turn, may be 918 translated into a cumulative ramping activity for each constituent movement by a 919 separate downstream neuronal mechanism during the preparation period (Cisek and 920 Kalaska 2005; Li et al. 2016). Analysis of the 'sequenceness' of the corresponding 921 neural patterns (Eichenlaub et al. 2020; Liu et al. 2019) during preparation should shed 922 light on the presence of preplay and its possible relationship to the competitive pre-923 activation of movements during planning (Kornysheva et al. 2019).

924 Characteristic differences in press error rate to movement probes were 925 revealed through faster rather than slower responses after the *Probe* cue (Figure 2c).

This suggests that the competitive pre-activation gradient established during the short phase of sequence retrieval and planning is driven by a rapid automatic process and is not a result of slow deliberation or higher-level decision making. Contrary to a prominent account of sequence learning (Krakauer and Mazzoni 2011; Wong and Krakauer 2019), we suggest that the reported behavioural differences in sequence press availability reflect mechanisms of rapid and automatic planning for the production of discrete motor sequences from memory.

933 Remarkably, longer preparation reinforced the competitive pre-activation 934 making responses to movement probes associated with later sequence positions even 935 slower and more inaccurate relative to those associated with earlier positions. This is counterintuitive in the context of single movement performance gains from longer 936 937 foreperiod durations (Niemi and Näätänen 1981). Here, a pure foreperiod effect would 938 dictate general benefits for RT and error rate with longer preparation durations 939 (Steinborn et al. 2008). In contrast, we found relative benefits and costs of the latter to 940 be position-dependent. The reported differences in movement availability became 941 more striking the longer time participants had to prepare, e.g., the error rate for probed 942 movements associated with later positions increased further with longer foreperiods -943 these movements became even harder to retrieve. The pre-activation gradient 944 expansion with longer preparation suggests a dynamic refinement of the plan for 945 sequence production during retrieval and planning. We propose that the primacy 946 gradient (Grossberg 1978a, 1978b) in the parallel planning layer of CQ models 947 expands dynamically during each sequence preparation phase enhancing the 948 organisation of sequential movements with preparation time.

949 Furthermore, participants exhibiting more pronounced differences in availability of movements associated with neighbouring sequence positions during planning 950 951 exhibited both faster initiation times and a more accurate temporal execution of the sequence after the Go cue, particularly when looking at position-dependent 952 953 differences in RT. These findings strengthen the interpretation that an ordered 954 competitive pre-activation of movements during planning pre-empts subsequent 955 fluency and temporal accuracy of the sequence (Kornysheva et al. 2019). The 956 individual differences in planning are likely driven by differences in sequence learning.

which are associated with an expansion of the "planning horizon" with practice (Arianiet al. 2020).

Yet, we did not replicate the association of the planning gradient with finger error 959 960 probability found in the latter study. This may be due to a smaller pool of timing and finger order sequences that the participants had to learn relative to the previous 961 962 paradigm, and the presence of only one finger order (paired with different sequence 963 timings) in Experiments 2 and 3. This facilitated finger accuracy to reach ceiling levels 964 in a substantial number of participants. Future experiments should resolve an 965 association with finger accuracy through the inclusion of a larger pool of trained 966 sequences to provoke more frequent finger errors. Alternatively, reaching, drawing or force production tasks would allow to quantify more fine-grained deviations from target 967 968 at overall high ordinal accuracy levels of sequence production.

969 In contrast to preparation duration, doubling the speed of sequence production 970 did not change the relative behavioural availability of sequential movements during 971 planning. This suggests invariance of the pre-activation gradient across sequences 972 produced at different time scales. This transfer across speed profiles is in line with the 973 presence of flexible motor timing and temporal scaling in dynamic neuronal 974 populations (Goudar and Buonomano 2018; Wang et al. 2018). Here the assumption 975 is that a separate neural process controls the speed of a sequence during execution, 976 e.g. through the strength of an external input to the network involved in the generation 977 of timed behaviour (Wang et al. 2018). We found that preparing a sequence of the 978 same length with an irregular compared to an isochronous interval structure was 979 associated with a slight tendency for a dampened CQ gradient during sequence 980 planning. However, this non-significant trend is unlikely to be the effect of temporal 981 grouping, as the irregular interval sequence was characterized by a significant 982 increase in temporal interval production error (Figure 5b, middle), associated with timing complexity - the sequencing of two different (non-isochronous) constituent 983 984 temporal intervals rather than just one (isochronous). Instead, we hypothesize that 985 longer preparation time (above 1500 ms) would have benefitted the participants and 986 enhanced the relative pre-activation gradient, in line with Experiment 1, facilitating the 987 formation of a more accurate plan for this more temporally complex sequence.

Our empirical data on the pre-ordering of sequential movements does not 988 989 support the integration of movement order with movement timing prior to sequence 990 execution. The weighting of the availability of each movement appears to be entirely 991 driven by its position in the planned sequence and correlated with the fluency of correct 992 sequence initiation. Given that participants could on average correctly modulate the 993 relative timing of the sequences, a separate preparation process for the speed and 994 timing of the respective sequence must be assumed. The latter may take place 995 concurrently or at different time points during preparation (Bortoletto et al. 2011; 996 Bortoletto and Cunnington 2010: Maslovat et al. 2018). In previous work, we proposed a drift-diffusion based model which contains input from separate modules that activate 997 998 movement order and timing (Kornysheva et al. 2013). This model was based on 999 behavioural sequence learning data demonstrating that sequence timing is encoded 1000 independently of the movement order, but requires multiplicative, rather than additive 1001 integration with each movement. This enables trained sequence timing to be 1002 transferred to new sequences, but only after the movement order has been acquired. 1003 reconciling previous experimental findings (Kornysheva and Diedrichsen 2014; 1004 O'Reilly et al. 2008; Shin and Ivry 2003; Ullén and Bengtsson 2003; Zeid and Bullock 1005 2019).

1006 Recently, Zeid and Bullock proposed how such plans may be generated in the 1007 context of CQ models (Zeid and Bullock 2019). The authors propose that two separate 1008 CQ modules could operate in parallel - one controlling the item order and the other 1009 controlling the sequence of inter-press-intervals that define a rhythmic pattern, 1010 including separate parallel planning and competitive choice layers. While this model 1011 is in line with neurophysiological and imaging evidence for a separate control of timing 1012 for sequence generation (Bengtsson et al. 2004, 2005; Crowe et al. 2014; Friston and 1013 Buzsáki 2016; Kornysheva and Diedrichsen 2014; Merchant et al. 2013), empirical 1014 support for timing being implemented via a CQ process for temporal intervals is still 1015 lacking. Behavioural paradigms are unlikely to be valuable in this context, as it is 1016 impossible to probe the planning of inter-press-interval sequences decoupled from the 1017 effector. However neurophysiological recordings in monkeys and humans may shed 1018 further light on the organisation of interval patterns prior to production: If temporal intervals in a sequence are competitively queued, we should expect neuronal 1019

1020 populations preferentially tuned to temporal intervals of different durations, e.g. as 1021 found in the medial premotor cortex (Crowe et al. 2014; Merchant et al. 2013), to be 1022 pre-activated in parallel during planning according to their respective position in the 1023 sequence, and transfer across effectors.

1024 Alternatively, timing of discrete movements in a sequence may be controlled 1025 during execution only through the acquired cyclical dynamics of neuronal population 1026 activity. Specifically, isochronous sequences involving the same movement have been 1027 associated with circular population trajectories where each interval cycle is shifted forward along a sequence position or "tapping manifold" resulting in a helical 1028 1029 population trajectory (Balasubramaniam et al. 2020; Russo et al. 2020). Here the interval duration has been linked to the amplitude size of the trajectory loops thus 1030 1031 controlling the speed of isochronous tapping sequences. The sequence position or 1032 "tapping manifold" may be the readout of a competitive queuing process and thus 1033 serve as a potential interface between position, interval, and movement identity. 1034 However, it remains unclear whether such a cyclical procession of population activity 1035 is also utilised for the production of sequences with non-isochronous intervals and 1036 sequences involving multiple movements.

1037

1038 Conclusions

In sum, our findings indicate that the behavioural availability of movements 1039 1040 during a brief period of retrieval and planning reflects the subsequent movement order, 1041 such that movements associated with later positions are less available for fast and 1042 accurate execution. Crucially, the competitive state of the movements appears to be 1043 invariant to the exact timing of the sequence. Instead, it is dynamically established 1044 during sequence planning and predicts the individual's subsequent sequence production fluency and accuracy. The current behavioural paradigm could provide a 1045 1046 straightforward and cost-effective way to assess the organisation of movements during 1047 sequence planning across trials in individual participants, in addition to 1048 neurophysiological approaches requiring access to neuroimaging, electrophysiology 1049 and computational resources for advanced neural pattern analysis (Averbeck et al. 1050 2002; Kornysheva et al. 2019). This behavioural readout of the state of movements 1051 before execution may serve to advance our understanding of the neural processes

associated with disorders affecting the fluent production of motor sequences, such as
stuttering, dyspraxia, and task-dependent dystonia (Craig-McQuaide et al. 2014;
Howell 2007; Ingham et al. 2018; Miller 1988; Sadnicka et al. 2018).

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1056 Supplemental Figure S1: https://doi.org/10.6084/m9.figshare.13688131

1057 Supplemental Figure S2: https://doi.org/10.6084/m9.figshare.13227953

1058 Supplemental Figure S3: https://doi.org/10.6084/m9.figshare.13168514

1059 Supplemental Figure S4: https://doi.org/10.6084/m9.figshare.13168628

1060 Supplemental Figure S5: https://doi.org/10.6084/m9.figshare.13168649

1061 Supplemental Figure S6: https://doi.org/10.6084/m9.figshare.13675330

1062 Supplemental Table S1: https://doi.org/10.6084/m9.figshare.13673605

1063 Supplemental Table S2: https://doi.org/10.6084/m9.figshare.13673668

1064 Supplemental Table S3: https://doi.org/10.6084/m9.figshare.13673734

1065 Supplemental Table S4: https://doi.org/10.6084/m9.figshare.13673800

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1301

1302 Figure legends

1303 Figure 1 | Design and experimental conditions. a. The first two days integrated the three training 1304 stages. Participants progressed from entirely instructed sequence production trials (stage A) to blocks 1305 of mixed trials (stage B) and, finally, to producing the target sequences from memory during the last 1306 stage of the training (stage C). All training stages incorporated a fixed percentage of Probe trials, 1307 randomized in each block, to ensure a degree of familiarity with single-press Probe cues. In the test 1308 phase (Day 3), participants underwent two refresher blocks (stage B) and, subsequently, an equal 1309 number of memory-guided Sequence trials and Probe trials (test). b. Test phase: After training, 1310 participants were prompted to produce 4-element finger sequences from memory following a Go cue. 1311 Each finger order or timing corresponded to a unique abstract visual Sequence cue presented for up to 1312 1500 ms before the Go cue (preparation period). Experiment 1 cued the production of sequences with 1313 two different finger orders and isochronous timing (slow). Here, we manipulated the duration of the 1314 preparation period (500, 1000, 1500 ms). In Experiments 2 and 3, the Sequence cues had a fixed 1315 preparation duration of 1500 ms and prompted the production of sequences with the same finger order 1316 but a different timing (slow, fast, irregular). In all three experiments, the target IPIs, illustrated in ms, 1317 were used to train participants to develop a relative timing proportionate to the target timing. Participants 1318 received visual feedback in each trial on the accuracy of the finger order and their timing. Points were 1319 based on finger press accuracy, initiation reaction time (RT), and temporal accuracy (cf. Materials and 1320 Methods). c. Test phase: In all experiments, we introduced Probe trials, in which, following the 1321 preparation period, the Go cue was replaced with a Probe cue. That prompted a particular finger digit 1322 to be pressed, corresponding to each sequence position or a control movement which did not feature 1323 in any sequence production. The Probe condition was used to obtain the RT and error rate for each 1324 position at the end of the preparation period. The participants received points for accurate presses and 1325 fast RTs.

1326 Figure 2 | Position-dependent movement availability during sequence planning. a. RTs for each 1327 probed sequence position relative to the first position. b. Press errors for each probed sequence position 1328 relative to the first position. (cf. raw RT and press error graphs in Supplemental Figure S2a, b). Both 1329 relative RT and press error were calculated from RTs and press error rates, respectively, obtained in 1330 *Probe* trials prompting the production of a movement associated with the 1st - 4th press position of the 1331 planned sequence (Experiments 1, 2 and 3) or a control movement not present in any sequence 1332 (Experiment 3). Black inset violin plots illustrate the position-dependent increases of raw RT and raw press error in the baseline condition (Dur: 1500 ms, T: slow), from 1st to 2nd, 2nd to 3rd, and 3rd to 4th 1333 1334 positions. Grey inset violin plots illustrate the difference between 4th position and control across 1335 sequence conditions, as indicated by the brackets. c. Relative press error in lower ('Fast RT') and upper 1336 ('Slow RT') RT quartiles. Error bars in line graphs represent standard errors. In inset violin plots, solid 1337 white lines represent the median, and lower and upper dashed white lines represent the 25th and 75th 1338 percentiles, respectively. Significance asterisks over the black inset violin plots indicate one-tailed 1339 increases (position-dependent increases in RT and error rate), whereas the asterisks over the grey

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- 1340 inset violin plots represent significance for a two-tailed test (increases or decreases in availability 1341 relative to control movement). $|*P \le 0.05| **P \le 0.01| ***P \le 0.001|$

1342Figure 3 | Pattern of press errors for probed movements associated with different sequence1343positions. Incorrect presses per probed position across experiments are shown in percent of all1344responses. | * P ≤ 0.05 | ** P ≤ 0.01 | *** P ≤ 0.001 |

1345 Figure 4 | Correlation of performance with position-dependent differences in movement 1346 availability during planning. The mean difference between adjacent positions (1st - 2nd, 2nd - 3rd, 1347 3rd - 4th) based on RTs and press errors relative to the first position (Probe trials) was taken as a proxy 1348 for the pre-activation gradient size during preparation, with steeper (larger) differences reflecting a more 1349 expanded gradient (cf. raw RT and error differences in Supplemental Figure S4). a. Correlations 1350 between relative position-dependent differences in RT in Probe trials and each of the performance 1351 measures (initiation RT, relative temporal error, and finger error). b. Correlations between relative 1352 position-dependent differences in error rate in Probe trials and each of the performance measures 1353 (initiation RT, relative temporal error, and finger error). Inset graphs in each panel illustrate relative 1354 position-dependent RT (a) and press error (b) increases during planning for participants with faster vs 1355 slower initiation RT and lower vs higher relative temporal error performance (median splits). Error bars 1356 represent standard errors. All correlations are one-tailed, in line with one-sided predictions regarding 1357 the beneficial effect of a differentiated pre-activation of sequence movements during planning.

1358 Figure 5 | Sequence production. a. Relative timing as a function of inter-press interval (IPI) production 1359 of a slow, twice as fast and an irregular sequence. Both the produced (solid lines) and target IPIs 1360 (dashed lines) were normalized across trials relative to the baseline condition (Dur: 1500 ms, T: slow). 1361 Error bars represent standard errors. b. Sequence initiation RT (Go cue to first press latency), relative 1362 temporal error, and finger error (proportion of trials with incorrect presses) in each experimental 1363 condition (preparation duration, Experiment 1; timing, Experiments 2 and 3). Solid white lines represent the median, and lower and upper dashed white lines represent the 25th and 75th percentiles, 1364 1365 respectively. | * P \leq 0.05 | ** P \leq 0.01 | *** P \leq 0.001 |

Figure 1



Preparation 500/1000/1500 ms

Figure 2





Pressed position

Figure 4



Figure 5

